The square negative correlation property on ℓ_p^n balls

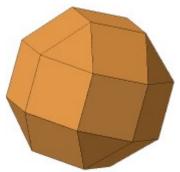
David Alonso Gutiérrez Joint work with Julio Bernués

Universidad de Zaragoza

7th of September, 2018

Convex bodies

• $K \subset \mathbb{R}^n$ is called a convex body if it is convex, compact and has non-empty interior.



A convex body $K \subseteq \mathbb{R}^n$ is isotropic if it has volume 1 and

- $\int_{\mathcal{K}} x dx = 0$ (centered at 0)
- $\int_{K} \langle x, \theta \rangle^2 dx = L_K^2 \quad \forall \theta \in S^{n-1}$.

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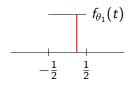
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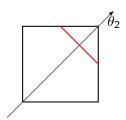
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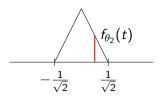




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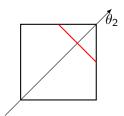


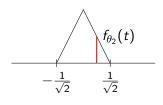


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Given K we consider a random vector X uniformly distributed in K and, for every $\theta \in S^{n-1}$, the real random variable $\langle X, \theta \rangle$ with density $f_{\theta}(t) = |K \cap (\theta^{\perp} + t\theta)|$.





K is isotropic if all the $\langle X, \theta \rangle$ are centered and have the same variance.



•
$$L_K \geq L_{B_2^n} = \frac{\Gamma\left(1+\frac{n}{2}\right)^{\frac{1}{n}}}{\pi\sqrt{n+2}}$$

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Hyperplane conjecture (Bourgain, 1990)

There exists an **absolute** constant C such that for every $K \subseteq \mathbb{R}^n$

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- $L_K \leq Cn^{\frac{1}{4}}$ (Klartag 2005, Lee-Vempala 2016)
- True for 1-unconditional bodies, polytopes with number of vertices proportional to the dimension, zonoids, unit balls of finite dimensional Schätten classes...

• X_1, \ldots, X_n independent random variables uniformly distributed in $[-\sqrt{3}, \sqrt{3}]$.

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• $G \sim \mathcal{N}(0,1)$

$$\lim_{n\to\infty} \sup_{t\in\mathbb{R}} \left| \mathbb{P}\left(\frac{X_1+\cdots+X_n}{\sqrt{n}} > t\right) - \mathbb{P}(G>t) \right| = 0$$



• $X = (X_1, \dots, X_n)$ random vector uniformly distributed in $K = \left[-\sqrt{3}, \sqrt{3}\right]^n$

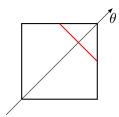
- $X = (X_1, \dots, X_n)$ random vector uniformly distributed in $K = \left[-\sqrt{3}, \sqrt{3} \right]^n$
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- X uniformly distributed in $K = L_{B_{\infty}^n}^{-1} B_{\infty}^n$, $\theta = \left(\frac{1}{\sqrt{n}}, \dots, \frac{1}{\sqrt{n}}\right)$, if n big

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Let X be a random vector uniformly distributed on $L_K^{-1}K$ with K isotropic. For how many directions $\theta \in S^{n-1}$ can we say

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Sudakov (1978), Diaconis-Freedman (1984), von Weizsäker (1997), Antilla-Ball-Perissinaki (2003).

Shown under a concentration hypothesis:

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Klartag (2007):

$$\mathbb{P}\left(\left|\frac{|X|}{\sqrt{n}} - 1\right| \ge \frac{c}{n^{\kappa}}\right) \le Ce^{-n^{\kappa}}$$

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Variance conjecture (Bobkov-Koldobsky, 2003)

There exists an **absolute** constant C such that for any isotropic log-concave random vector

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- $\operatorname{Var}|X|^2 \leq Cn^{\frac{1}{2}}L_K^2\mathbb{E}|X|^2$. (Lee-Vempala, 2016)
- True for uniformly distributed random vectors on Bⁿ_p
 (Antilla-Ball-Perissinaki, 2003), Orlicz balls (Wojtaszczyk, 2007),
 unconditional bodies (Klartag, 2012)



Definition

A centered log-concave random vector $X \in \mathbb{R}^n$ satisfies the square negative correlation property with respect to the orthonormal basis $\{\eta_i\}_{i=1}^n$ if for every $i \neq j$

$$\mathbb{E}\langle X,\eta_i\rangle^2\langle X,\eta_j\rangle^2 - \mathbb{E}\langle X,\eta_i\rangle^2\mathbb{E}\langle X,\eta_j\rangle^2 \leq 0$$

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- X uniformly distributed on $P_{\theta^{\perp}}B_{\infty}^{n}$, any $\{\eta_{i}\}_{i=1}^{n}$. A., Bastero (2013).

• X uniformly distributed on K isotropic, $\{\eta_i\}_{i=1}^n$

$$\begin{aligned} \operatorname{Var}|X|^2 &= \sum_{i=1}^n (\mathbb{E}\langle X, \eta_i \rangle^4 - (\mathbb{E}\langle X, \eta_i \rangle^2)^2) \\ &+ \sum_{i \neq j} (\mathbb{E}\langle X, \eta_i \rangle^2 \langle X, \eta_j \rangle^2 - \mathbb{E}\langle X, \eta_i \rangle^2 \mathbb{E}\langle X, \eta_j \rangle^2) \end{aligned}$$

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• If X satisfies the square negative correlation property with respect to $\{\eta_i\}_{i=1}^n$

$$|\operatorname{Var}|X|^2 \le \sum_{i=1}^n (\mathbb{E}\langle X, \eta_i \rangle^4 - (\mathbb{E}\langle X, \eta_i \rangle^2)^2)$$

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$$|\operatorname{Var}|X|^{2} \leq \sum_{i=1}^{n} (\mathbb{E}\langle X, \eta_{i} \rangle^{4} - (\mathbb{E}\langle X, \eta_{i} \rangle^{2})^{2})$$

$$\leq C \sum_{i=1}^{n} (\mathbb{E}\langle X, \eta_{i} \rangle^{2})^{2}$$

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- A random vector uniformly distributed on B_p^n satisfies the SNCP with respect to $\{e_i\}_{i=1}^n$. What about other basis?

Proposition (A., Bernués (2018)

Let X be a random vector uniformly distributed on a 1-symmetric convex body, $\xi_1 = \frac{e_1 + e_2}{\sqrt{2}}$, $\xi_2 = \frac{e_1 - e_2}{\sqrt{2}}$, and $f: S^{n-1} \times S^{n-1} \to \mathbb{R}$

$$f(\eta_1, \eta_2) = \mathbb{E}\langle X, \eta_1 \rangle^2 \langle X, \eta_2 \rangle^2 - \mathbb{E}\langle X, \eta_1 \rangle^2 \mathbb{E}\langle X, \eta_2 \rangle^2.$$

Then, if $\langle \eta_1, \eta_2 \rangle = 0$

$$f(\eta_1,\eta_2) = f(e_1,e_2) + 2(f(\xi_1,\xi_2) - f(e_1,e_2)) \sum_{i=1}^{n} \eta_1(i)^2 \eta_2(i)^2.$$

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$$0 \le \sum_{i=1}^n \eta_1(i)^2 \eta_2(i)^2 \le \frac{1}{2}$$
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$$f(\xi_1, \xi_2) - f(e_1, e_2) = \frac{\Gamma\left(1 + \frac{n-1}{p}\right)\Gamma\left(\frac{3}{p}\right)^2}{2\Gamma\left(1 + \frac{n+3}{p}\right)\Gamma\left(\frac{1}{p}\right)^2} \left(F\left(\frac{1}{p}\right) - 3\right)$$

$$F(x) = \frac{\Gamma(5x)\Gamma(x)}{\Gamma(3x)^2}.$$

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• F is strictly increasing in (0,1] and $F\left(\frac{1}{2}\right)=3$.



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Let X be a random vector uniformly distributed on B_p^n , $\xi_1 = \frac{e_1 + e_2}{\sqrt{2}}$, $\xi_2 = \frac{e_1 - e_2}{\sqrt{2}}$. If $\langle \eta_1, \eta_2 \rangle = 0$

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$$\bullet \ f(e_1, e_2) = \frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(\frac{3}{p}\right)^2}{\Gamma\left(1 + \frac{n+4}{p}\right)\Gamma\left(\frac{1}{p}\right)^2} \left(1 - \frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(1 + \frac{n+4}{p}\right)}{\Gamma\left(1 + \frac{n+2}{p}\right)^2}\right)$$

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•
$$f(e_1, e_2) = \frac{\Gamma\left(1 + \frac{n}{\rho}\right)\Gamma\left(\frac{3}{\rho}\right)^2}{\Gamma\left(1 + \frac{n+4}{\rho}\right)\Gamma\left(\frac{1}{\rho}\right)^2} \left(1 - \frac{\Gamma\left(1 + \frac{n}{\rho}\right)\Gamma\left(1 + \frac{n+4}{\rho}\right)}{\Gamma\left(1 + \frac{n+2}{\rho}\right)^2}\right)$$

• Since $\log \Gamma(x)$ is strictly convex $f(e_1, e_2) < 0$.

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- Since $\log \Gamma(x)$ is strictly convex $f(e_1, e_2) < 0$.
- $f(\eta_1, \eta_2) < 0$ for every η_1, η_2 with $\langle \eta_1, \eta_2 \rangle = 0$.

•
$$f(\xi_1, \xi_2) = \frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(\frac{3}{p}\right)^2 \left(F\left(\frac{1}{p}\right) - 1 - 2\frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(1 + \frac{n+4}{p}\right)}{\Gamma\left(1 + \frac{n+2}{p}\right)^2}\right)}{2\Gamma\left(1 + \frac{n+4}{p}\right)\Gamma\left(\frac{1}{p}\right)^2}$$

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• If $1 \le p < 2$, $F\left(\frac{1}{p}\right) > 3$

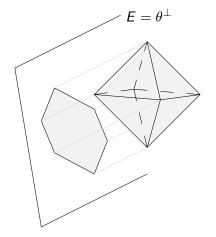
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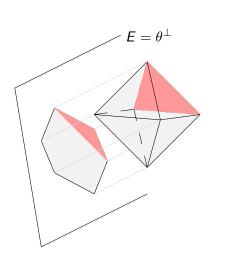
$$\bullet \ f(\xi_1, \xi_2) = \frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(\frac{3}{p}\right)^2 \left(F\left(\frac{1}{p}\right) - 1 - 2\frac{\Gamma\left(1 + \frac{n}{p}\right)\Gamma\left(1 + \frac{n+4}{p}\right)}{\Gamma\left(1 + \frac{n+2}{p}\right)^2}\right)}{2\Gamma\left(1 + \frac{n+4}{p}\right)\Gamma\left(\frac{1}{p}\right)^2}$$

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- If $1 \le p < 2$ there exists n_0 such that if $n \ge n_0$ $f(\xi_1, \xi_2) > 0$.

$$P_{\theta^\perp}K$$

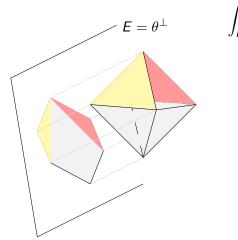


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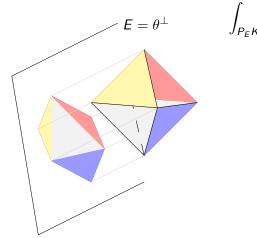
$$\int_{P_EK} f(x) dx =$$

$$P_{\theta^{\perp}}K$$



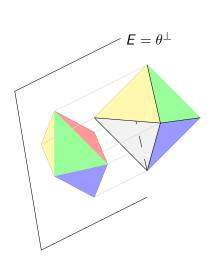
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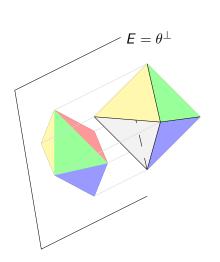
$$\int_{P_EK} f(x) dx =$$

$$P_{\theta^{\perp}}K = \bigcup_{F \in \mathcal{F}_{n-1}(K): \langle \nu(F), \theta \rangle \geq 0} P_E F$$



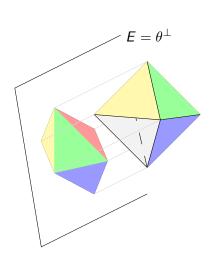
$$\int_{P_EK} f(x)dx = \sum_{F \in \mathcal{F}_{n-1}(K) \atop \langle \nu(F), \theta \rangle \geq 0} \int_{P_EF} f(x)dx$$

$$P_{\theta^{\perp}}K = \bigcup_{F \in \mathcal{F}_{n-1}(K): \langle \nu(F), \theta \rangle \ge 0} P_E F = \bigcup_{F \in \mathcal{F}_{n-1}(K): \langle \nu(F), \theta \rangle \le 0} P_E F$$



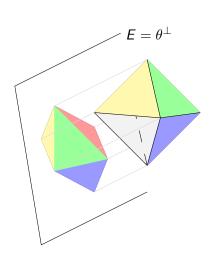
$$\int_{P_EK} f(x)dx = \sum_{\substack{F \in \mathcal{F}_{n-1}(K) \\ \langle \nu(F), \theta \rangle \leq 0}} \int_{P_EF} f(x)dx$$

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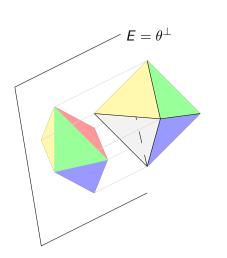
$$\int_{P_EK} f(x)dx = \frac{1}{2} \sum_{F \in \mathcal{F}_{n-1}(K)} \int_{P_EF} f(x)dx$$

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$$= \frac{1}{2} \int_{\partial K} f(P_E y) |\langle \nu(y), \theta \rangle| d\sigma_K(y)$$

 $d\sigma_K$ Hausdorff measure on ∂K

Random vectors on hyperplane projections of B_p^n

ullet By Cauchy's formula, if X is uniformly distributed on $P_{ heta^{\perp}}(B_p^n)$

$$\mathbb{E}f(X) = \frac{\int_{\partial B_{\rho}^{n}} f(P_{\theta^{\perp}}(x)) \frac{|\langle \nabla \| \cdot \|_{\rho}(x), \theta \rangle|}{|\nabla \| \cdot \|_{\rho}(x)|} d\sigma_{\rho}^{n}(x)}{\int_{\partial B_{\rho}^{n}} \frac{|\langle \nabla \| \cdot \|_{\rho}(x), \theta \rangle|}{|\nabla \| \cdot \|_{\rho}(x)|} d\sigma_{\rho}^{n}(x)}.$$

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• Since $d\sigma_p^n(x) = n|B_p^n||\nabla|| \cdot ||_p(x)|d\mu_p^n(x)$ (Naor, Romik, 2003)

$$\mathbb{E}f(X) = \frac{\int_{\partial B_p^n} f(P_{\theta^{\perp}}(x)) |\langle \nabla \| \cdot \|_p(x), \theta \rangle |d\mu_p^n(x)}{\int_{\partial B_p^n} |\langle \nabla \| \cdot \|_p(x), \theta \rangle |d\mu_p^n(x)}.$$

Probabilistic representation of $d\mu_p^n$ (Schechtman-Zinn (1990)

Let g_1, \ldots, g_n be independent copies of a random variable with density

$$\frac{e^{-|t|^p}}{2\Gamma\left(1+\frac{1}{p}\right)}$$

and

$$S = \left(\sum_{i=1}^{n} |g_i|^p\right)^{\frac{1}{p}}.$$

Then

- The random vector $\frac{g}{S} = (\frac{g_1}{S}, \dots, \frac{g_n}{S})$ and the random variable S are independent.
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If X is uniformly distributed on $P_{\theta^{\perp}}(B_p^n)$

$$\mathbb{E}f(X) = \frac{\mathbb{E}f\left(P_H\left(\frac{g}{S}\right)\right)\phi}{\mathbb{E}\phi}, \qquad \phi = \left|\sum_{i=1}^n \frac{|g_i|^{p-1}}{S^{p-1}} \operatorname{sgn}(g_i)\theta_i\right|$$



The square negative correlation property on hyperplane projections of B_p^n

$$\theta_0 = \left(\frac{1}{\sqrt{n}}, \dots, \frac{1}{\sqrt{n}}\right), P_{\theta_0^{\perp}}(B_p^n)$$
 is isotropic.

Theorem (A., Bernués, 2018)

 $f: S_{\theta^{\perp}} \times S_{\theta^{\perp}} \to \mathbb{R}$

Let X be a random vector uniformly distributed on $P_{\theta_0^{\perp}}(B_p^n)$, $\xi_1 = \frac{e_1 - e_2 + e_3 - e_4}{2}$, $\xi_2 = \frac{e_1 - e_2 - e_3 + e_4}{2}$, $\overline{\xi}_1 = \frac{e_1 - e_2}{\sqrt{2}}$, $\overline{\xi}_2 = \frac{e_3 - e_4}{\sqrt{2}}$ and

 $f(\eta_1, \eta_2) = \mathbb{E}\langle X, \eta_1 \rangle^2 \langle X, \eta_2 \rangle^2 - \mathbb{E}\langle X, \eta_1 \rangle^2 \mathbb{E}\langle X, \eta_2 \rangle^2.$

 $f(\xi_1, \xi_2) < f(\eta_1, \eta_2) < f(\overline{\xi}_1, \overline{\xi}_2).$

If $p \ge 2$ there exists $n_0(p)$ such that if $n \ge n_0$ and $\langle \eta_1, \eta_2 \rangle = 0$

If
$$1 \le p \le 2$$
 there exists $n_1(p)$ such that if $n \ge n_1$ and $\langle \eta_1, \eta_2 \rangle = 0$
$$f(\overline{\xi}_1, \overline{\xi}_2) \le f(\eta_1, \eta_2) \le f(\xi_1, \xi_2).$$

The square negative correlation property on hyperplane projections of B_p^n

Corollary (A., Bernués, 2018)

Let X be a random vector uniformly distributed on $P_{\theta_0^{\perp}}(B_p^n)$. There exists $n_0(p)$ such that for all $n \geq n_0 X$ satisfies the SNCP with respect to every orthonormal basis in θ_0^{\perp} .